

MACHINE LEARNING BEYOND THE DATA RANGE: EXTREME QUANTILE REGRESSION

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SEMINAR IN MATHEMATICS, PHYSICS & MACHINE LEARNING
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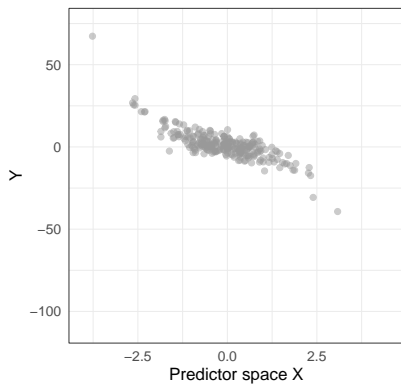


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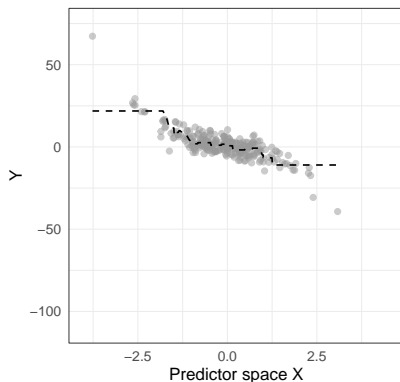


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Out-of-sample prediction: domain generalization

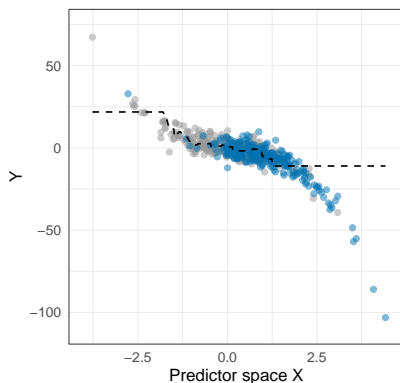


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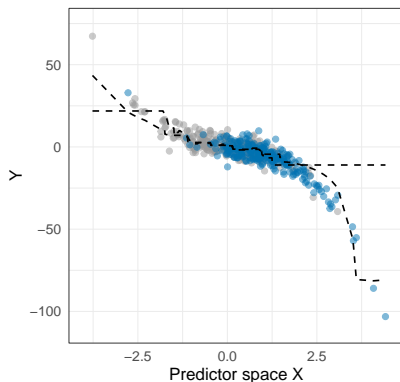
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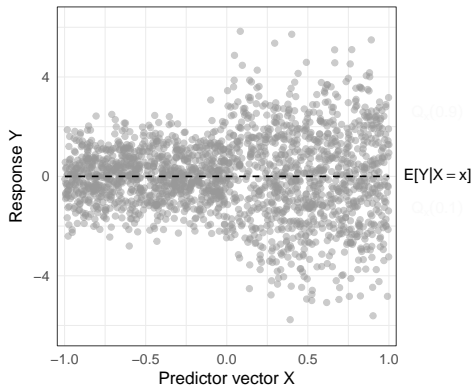
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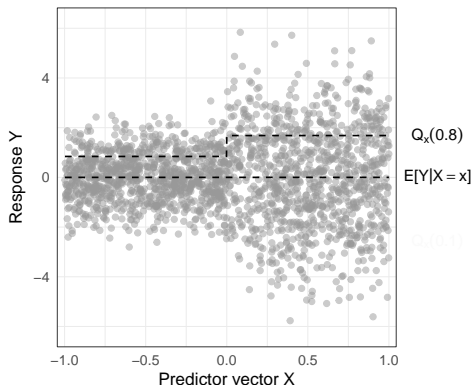


- ▶ Prediction of condition mean $\mathbb{E}[Y \mid X = x]$
- ▶ Machine learning methods typically perform poorly outside of the range of training distribution
- ▶ If test distribution is different, domain generalization is needed
- ▶ We work on methods based on **extreme value theory**

Out-of-sample prediction: extreme quantile regression

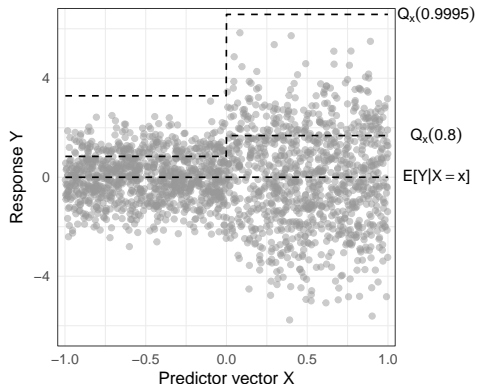


Out-of-sample prediction: extreme quantile regression



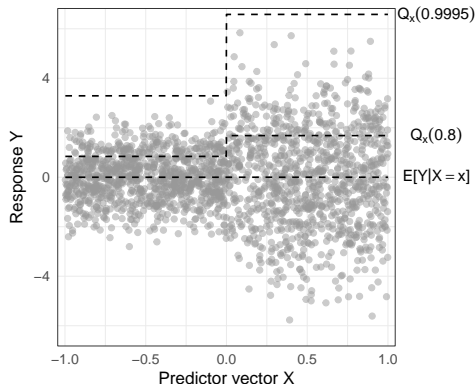
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- ▶ If τ is close to 1, we speak of **extreme quantile regression**
- ▶ Again, classical machine learning methods perform poorly

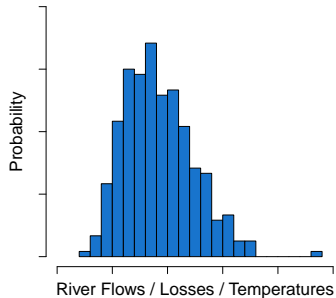
Extreme Value Theory and Statistics

- ▶ Analysis of **rare phenomena** with small probabilities
- ▶ Impact on **various risks** (health, environment, economy,...)



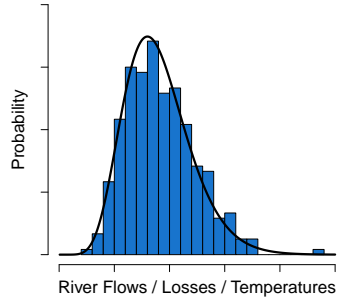
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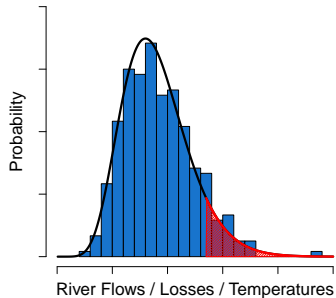
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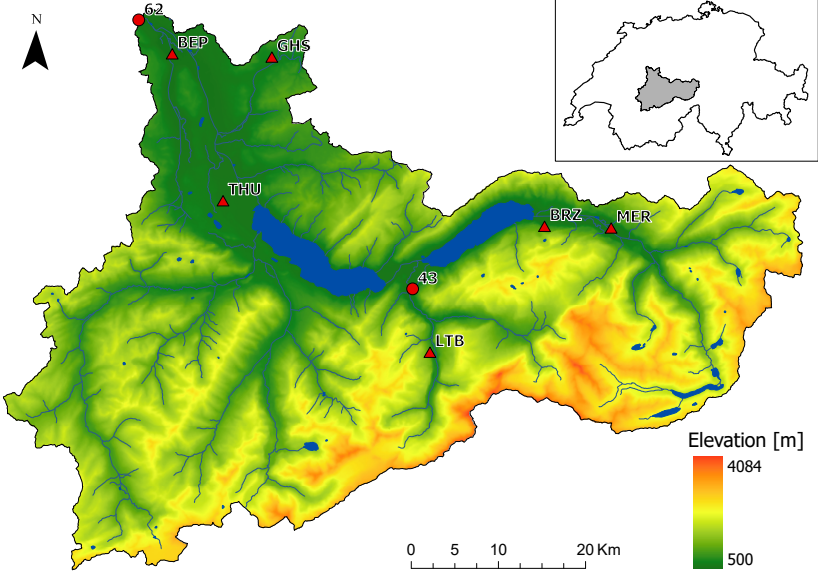


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Aare River in Bern: Catchment



Data

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- ▶ Training and validation data: 1930–1958 (10,349 obs.).
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As regression problem

- ▶ Y is daily discharge at Bern station.
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Extreme quantile regression

- ▶ For i.i.d. data $(\mathbf{X}_1, Y_1), \dots, (\mathbf{X}_n, Y_n)$ where $\mathbf{X}_i \in \mathbb{R}^p$ and $Y_i \in \mathbb{R}$, the goal is to predict the conditional quantile at level $\tau \in (0, 1)$

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- ▶ There are different scenarios depending on the quantile level $\tau = \tau_n$:
 - ▶ $\tau_n \equiv \tau_0 < 1$ (classical case)
 - ▶ $\tau_n \rightarrow 1$, and $n(1 - \tau_n) \rightarrow \infty$ (intermediate case)
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- ▶ Methods from extreme value theory are not flexible enough [Chernozhukov, 2005, Chavez-Demoulin and Davison, 2005] or do not generalize well into higher dimensions [Daouia et al., 2011].

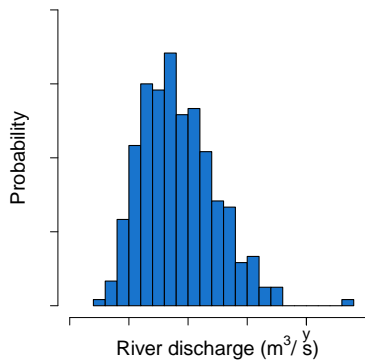
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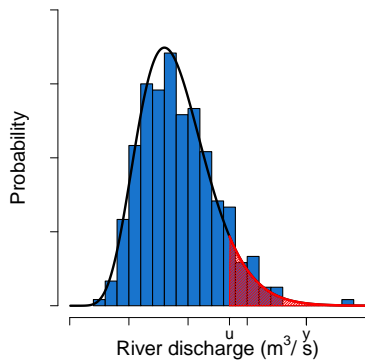
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- ▶ **Goal:** Develop a new method for **extreme quantile regression** that works well with high-dimensional and complex data.

Generalized Pareto distribution



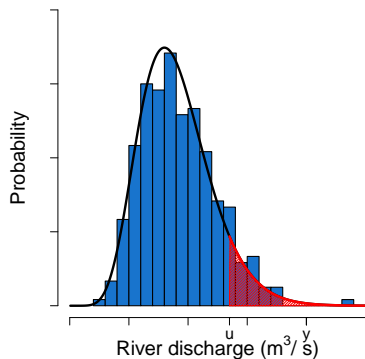
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$$\mathbb{P}(Y > y) = \mathbb{P}(Y > u) \times \mathbb{P}(Y > y \mid Y > u)$$

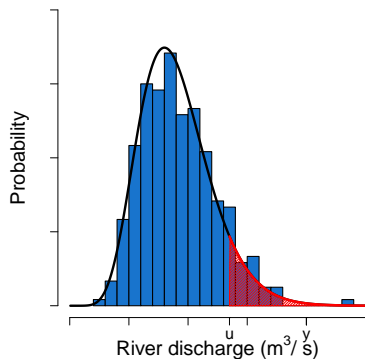
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$$\begin{aligned}\mathbb{P}(Y > y) &= \mathbb{P}(Y > u) \times \mathbb{P}(Y > y \mid Y > u) \\ &\approx \mathbb{P}(Y > u) \times (1 - H_{\sigma, \gamma}(y - u))\end{aligned}$$

where $H_{\sigma, \gamma}$ is the cdf of the GPD with **scale** and **shape** $\sigma > 0$ and $\gamma \in \mathbb{R}$.

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Estimation

- ▶ Consider i.i.d. data Y_1, \dots, Y_n and estimate empirically the quantile $u = \hat{Q}(\tau_0)$ for an **intermediate quantile** level $\tau_0 < 1$.
- ▶ Define the **exceedances** above the threshold as

$$Z_i = \left(Y_i - \hat{Q}(\tau_0) \right)_+ .$$

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- ▶ The **likelihood** of the GPD model with parameters $\theta = (\sigma, \gamma)$ is

$$\ell_{Z_i}(\theta) = - \left[(1 + 1/\gamma) \log \left(1 + \gamma \frac{Z_i}{\sigma} \right) + \log \sigma \right] \mathbb{I}_{Z_i > 0} .$$

Estimate the parameters by maximum likelihood

$$\hat{\theta} = \arg \max_{\theta} \sum_{i=1}^n \ell_{Z_i}(\theta) .$$

Extreme quantile estimation

- ▶ Inverting the cdf $H_{\hat{\sigma}, \hat{\gamma}}$ of the GPD provides an approximation of the quantile $Q(\tau) = F_Y^{-1}(\tau)$ for probability level $\tau > \tau_0$ by

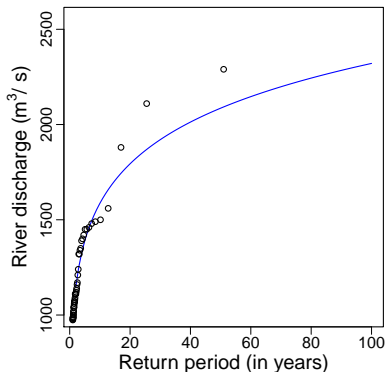
$$\hat{Q}(\tau) = \hat{Q}(\tau_0) + \hat{\sigma} \frac{\left(\frac{1-\tau}{1-\tau_0}\right)^{-\hat{\gamma}} - 1}{\hat{\gamma}}.$$

Extreme quantile estimation

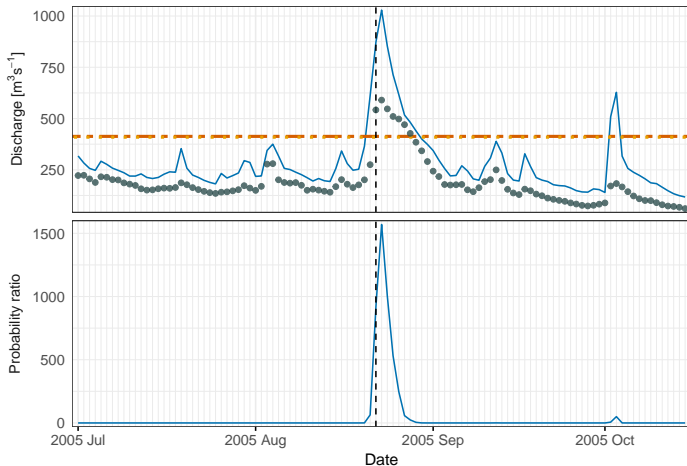
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- ▶ For independent data, we can compute a T -year event by $\hat{Q}(1 - 1/(n_Y T))$, where n_Y is the number of **observations per year**.



During the 2005 flood



- ▶ Top: Daily observations during the 2005 flood in Bern together with 100-year return level estimate. Vertical dashed line is the first exceedance of this level.

Forest-based quantile regression

- ▶ For $\tau \in (0, 1)$ and $x \in [-1, 1]^p$, the quantile regression function is defined as

$$Q_{\mathbf{x}}(\tau) := \arg \min_q \mathbb{E}[\rho_{\tau}(Y - q) \mid \mathbf{X} = \mathbf{x}],$$

where $\rho_{\tau}(c) = c(\tau - \mathbb{1}\{c < 0\})$ is the quantile loss function.

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- ▶ Recently, [Meinshausen, 2006] and [Athey et al., 2019] proposed to estimate $Q_{\mathbf{x}}(\tau)$ by

$$\hat{Q}_{\mathbf{x}}(\tau) = \arg \min_q \sum_{i=1}^n w_n(\mathbf{x}, X_i) \rho_{\tau}(Y_i - q),$$

where $(x, y) \mapsto w_n(\mathbf{x}, y)$ is a localizing weight function learned with a **random forest**.

Extreme quantile regression

Given: Independent data $(\mathbf{X}_1, Y_1), \dots, (\mathbf{X}_n, Y_n)$ of response $Y \in \mathbb{R}$ and covariates/predictors vector $\mathbf{X} \in \mathbb{R}^d$.

Goal:

- ▶ Predict extreme conditional quantiles of Y given $\mathbf{X} = \mathbf{x}$:

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- ▶ Predict exceedance probability of Y over high threshold Q (e.g., 100-year level) given $\mathbf{X} = \mathbf{x}$:

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Examples:

- ▶ Bern river discharge:
 - ▶ Y is daily discharge at Bern station.
 - ▶ \mathbf{X} can contain: discharge from previous days at same and other stations; precipitation from close-by stations; other climatological variables.
- ▶ Risk of heat waves:
 - ▶ Y daily temperature measurement at some location.
 - ▶ \mathbf{X} can contain: altitude; day of the year; other variables on land use, climate, etc.

Extreme quantile regression

- ▶ Assume the **GPD model**

$$(Y - \hat{Q}(\tau_0) \mid Y > Q(\tau_0), \mathbf{X} = \mathbf{x}) \sim H_{\sigma, \gamma},$$

where τ_0 is an **intermediate** quantile level, and $\hat{Q}_{\mathbf{x}}(\tau_0)$ is an estimate of the conditional τ_0 quantile of $Y \mid \mathbf{X} = \mathbf{x}$

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- ▶ For an **extreme level** $\tau > \tau_0$ we can estimate

$$\hat{Q}_{\mathbf{x}}(\tau) = \hat{Q}_{\mathbf{x}}(\tau_0) + \hat{\sigma}(\mathbf{x}) \frac{\left(\frac{1-\tau}{1-\tau_0}\right)^{-\hat{\gamma}(\mathbf{x})} - 1}{\hat{\gamma}(\mathbf{x})},$$

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- ▶ The triple $(\hat{Q}_{\mathbf{x}}(\tau_0), \hat{\sigma}(\mathbf{x}), \hat{\gamma}(\mathbf{x}))$ provides a model for the tail of $Y \mid \mathbf{X} = \mathbf{x}$.

Extreme quantile regression

Two methods to estimate the GPD parameters $\hat{\theta}(\mathbf{x}) = (\hat{\sigma}(\mathbf{x}), \hat{\gamma}(\mathbf{x}))$, both maximize a **localized likelihood**:

$$\sum_{i=1}^n w_n(\mathbf{x}, X_i) \ell_{(\sigma, \gamma)}(Z_i)$$

$\ell_{(\sigma, \gamma)}(Z_i)$ is the GPD **log-likelihood**, with exceedances $Z_i = (Y_i - \hat{Q}_{\mathbf{x}}(\tau_0))_+$.

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- ▶ **Extremal gradient boosting (GBEX)**: The weights $w_n(\mathbf{x}, X_i)$ are obtained through gradient boosting.



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Gradient boosting for extreme quantile regression.

<https://arxiv.org/abs/2103.00808>

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Gnecco, N., Terefe, E.M., and Engelke, S. (2022).

Extremal Random Forests.

<https://arxiv.org/abs/2201.12865>

Extremal Random Forest (ERF)

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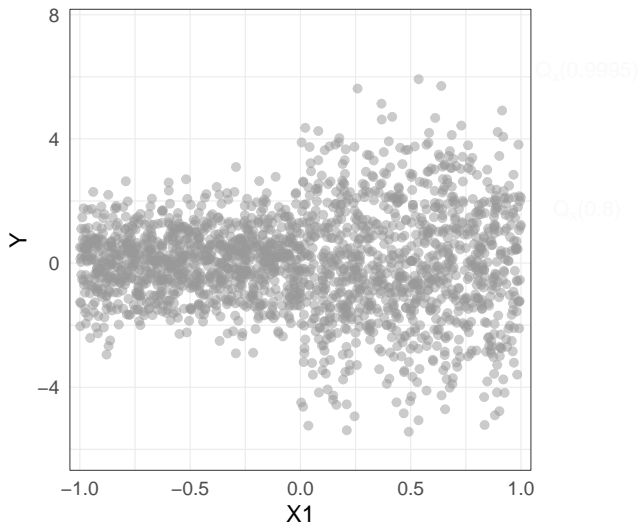
Extremal Random Forest (ERF)

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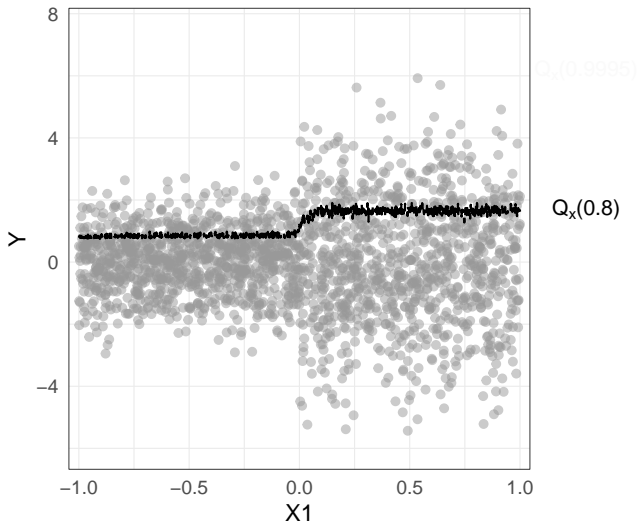
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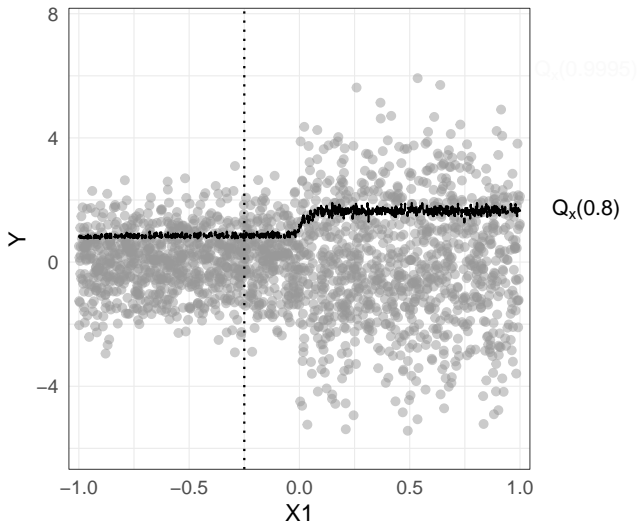
Extremal Random Forest (ERF)



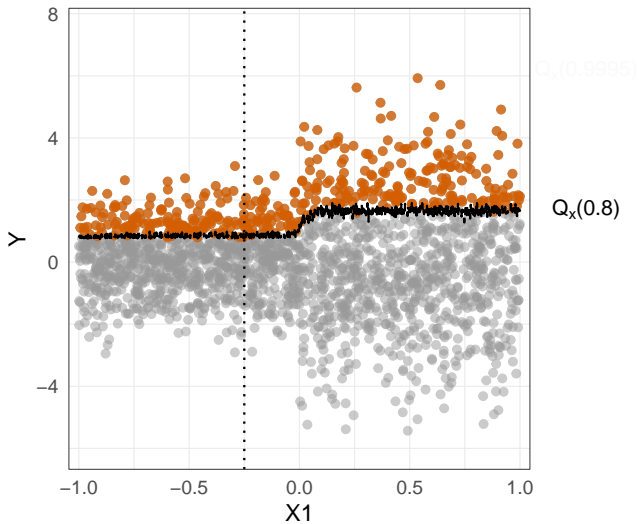
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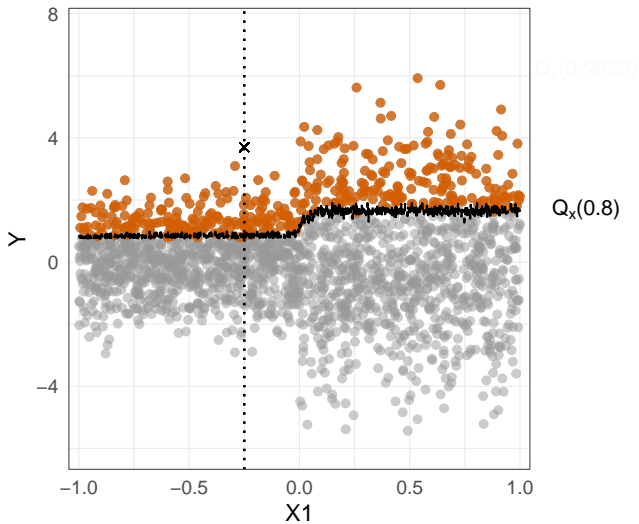
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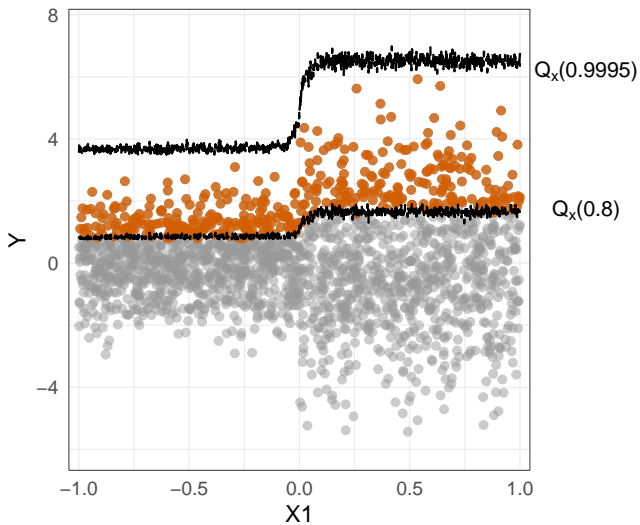
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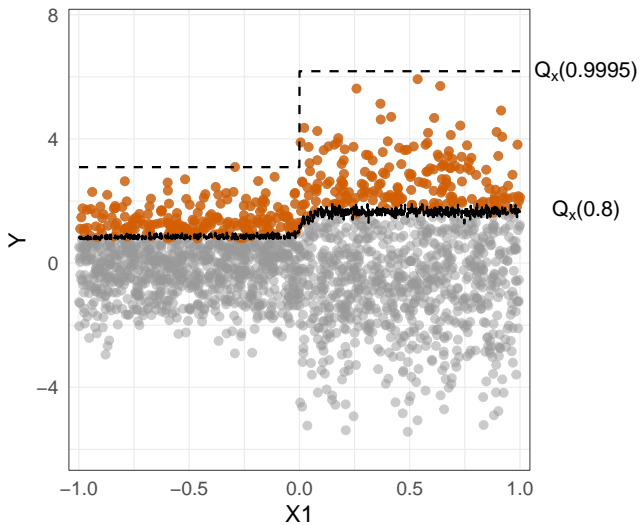
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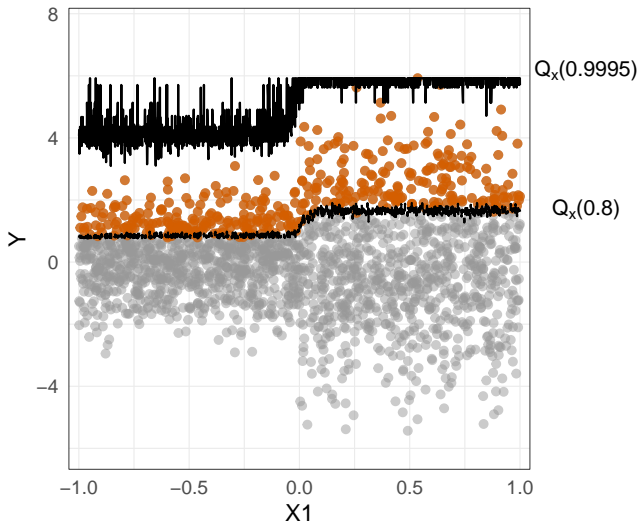
Extremal Random Forest (ERF)



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Consistency

- ▶ Let $x \in [-1, 1]^p$ and $\theta := (\sigma, \gamma)$. Want to show that

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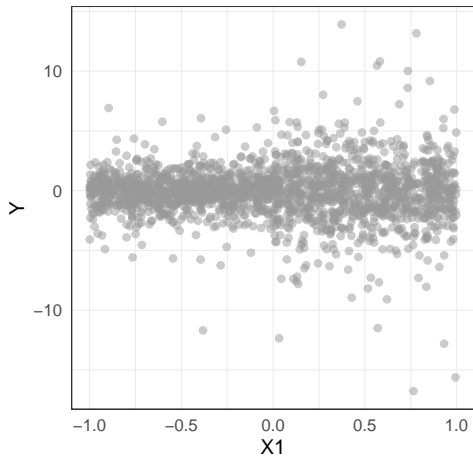
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- ▶ Regularity conditions from [Athey et al., 2019] do not hold in our setting.
- ▶ Under some assumptions, ERF estimates are **consistent**

$$\hat{\theta}(x) \xrightarrow{\mathbb{P}} \theta(x), \quad \text{for all } x \in [-1, 1]^p.$$

Simulation Study I



Simulation Study I

- ▶ Sample $n = 2000$ iid copies of (X, Y) from

$$\begin{cases} \mathbf{X} \sim U([-1, 1]^p), \\ (Y | \mathbf{X} = \mathbf{x}) \sim s(\mathbf{x})T_4, \end{cases}$$

where $s(x) = 1 + \mathbb{1}\{x_1 > 0\}$ and $\gamma(x) = 1/4$.

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- ▶ Compare ERF and GBEX with QRF [Meinshausen, 2006], GRF [Athey et al., 2019] Extreme GAM [Youngman, 2019], [Taillardat et al., 2019].

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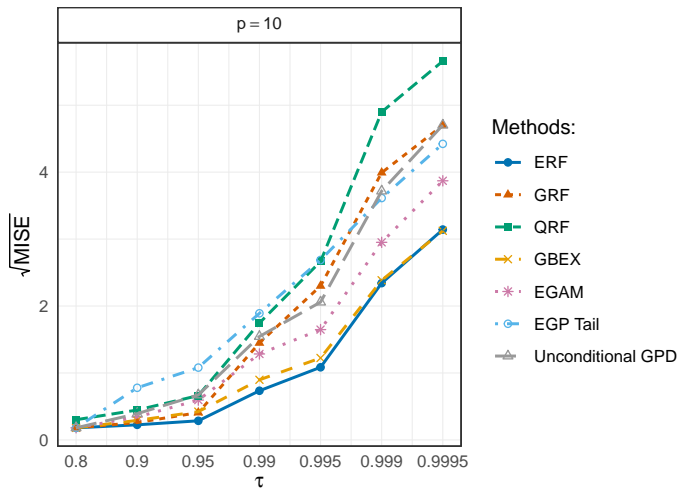
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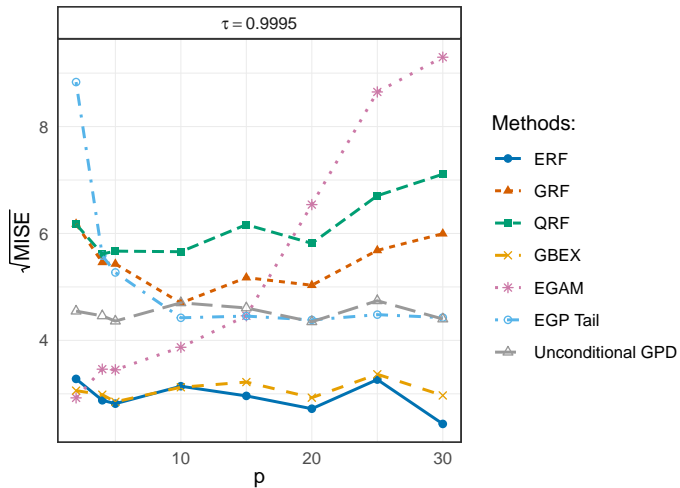
- ▶ Compare ERF and GBEX with QRF [Meinshausen, 2006], GRF [Athey et al., 2019] Extreme GAM [Youngman, 2019], [Taillardat et al., 2019].
- ▶ On a test data set $\{\mathbf{x}_i\}_{i=1}^{n'}$, evaluate the integrated squared error (ISE)

$$\text{ISE} = \frac{1}{n'} \sum_{i=1}^{n'} \left(\hat{Q}_{\mathbf{x}_i}(\tau) - Q_{\mathbf{x}_i}(\tau) \right)^2.$$

Simulation Study I



Simulation Study I



Simulation Study II

- ▶ Sample $n = 5000$ iid copies of (X, Y) from

$$\begin{cases} \mathbf{X} \sim U([-1, 1]^p), \\ (Y \mid \mathbf{X} = \mathbf{x}) \sim s_j(\mathbf{x})T_{\nu(\mathbf{x})}, \quad j = 1, 2, 3. \end{cases}$$

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- ▶ $\nu(x) := 3 + 3[1 + \tanh(-2X_1)]$ where $\gamma(x) = 1/\nu(x)$.

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Simulation Study II

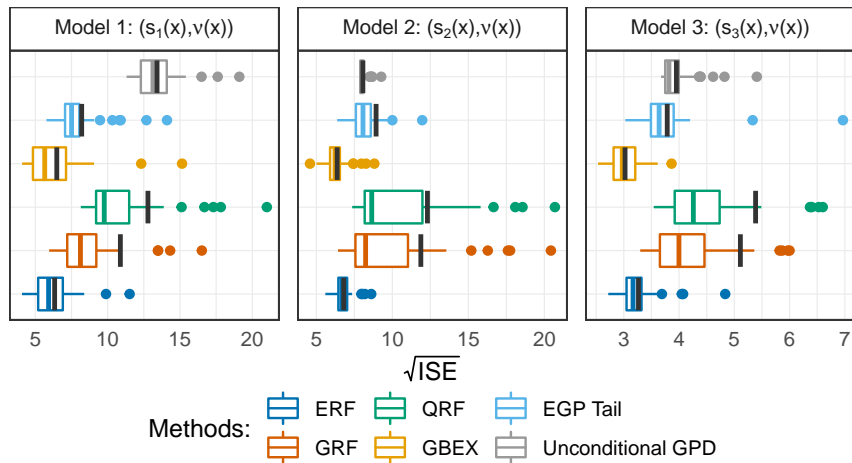
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- ▶ $s_2(x) := 4 - (X_1^2 + 2X_2^2)$.
- ▶ $s_3(x) := 1 + 2\pi\phi(2X_1, 2X_2)$, where $\phi(X_1, X_2)$ is a centered bivariate Gaussian with unit variance and correlation equal to $3/4$.

Simulation Study II

$n = 5000$, $p = 10$, $\tau = 0.9995$



Extreme quantile regression

- ▶ Suppose the data $(\mathbf{X}_1, Y_1), \dots, (\mathbf{X}_n, Y_n)$ are NOT i.i.d., but have a time series structure
- ▶ Then we can use **recurrent neural networks** to model the GPD parameters $(\sigma(\mathbf{x}), \gamma(\mathbf{x}))$
- ▶ **Extreme quantile regression neural networks (EQRN):**



Pasche, O.C. and Engelke, S. (2022).

Neural Networks for Extreme Quantile Regression with an Application to Forecasting of Flood Risk.

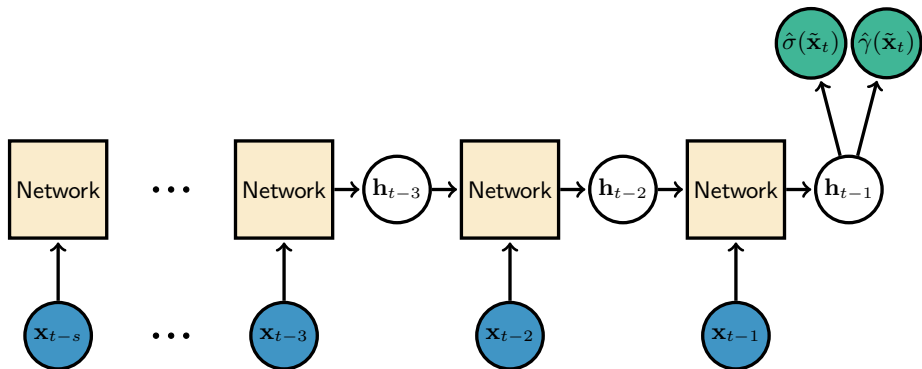
<https://arxiv.org/abs/2208.07590>

Extreme quantile regression neural networks (EQRN)

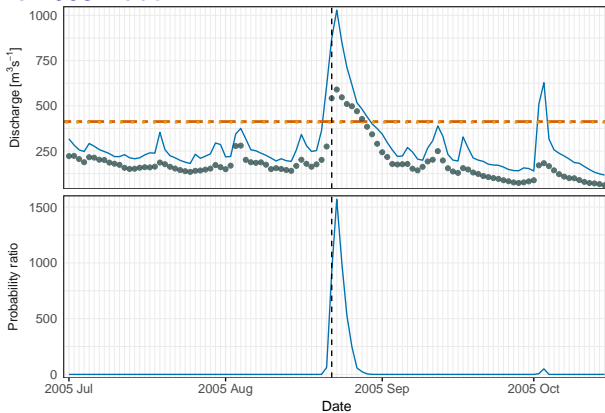
- ▶ If there is sequential dependence as in **time series**, then this structure can be used in **recurrent neural networks**.
- ▶ Predict quantiles of Y_t (discharge at time t) using past observations

$$\mathbf{X} = (Y_{t-1}, Y_{t-2}, \dots, X_{t-1}^1, X_{t-2}^1, \dots)$$

from response on other covariates X^1, X^2, \dots (e.g., precipitation at locations 1, 2, etc.).

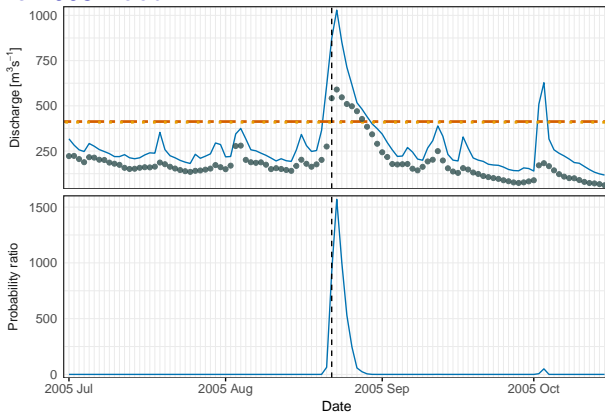


Results for the 2005 flood



- ▶ Top: Blue line is the one-day-ahead forecasted (conditional) 100-year return level $\hat{Q}_{\mathbf{x}}^{100}$.

Results for the 2005 flood



- ▶ Top: Blue line is the one-day-ahead forecasted (conditional) 100-year return level \hat{Q}_x^{100} .
- ▶ Bottom: Blue line the the ratio of conditional exceedance probability compared to unconditional estimate

$$\frac{\hat{\mathbb{P}}(Y > \hat{Q}_x^{100} \mid \mathbf{X} = \mathbf{x})}{\hat{\mathbb{P}}(Y > \hat{Q}^{100})}$$

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Youngman, B. D. (2019). Generalized additive models for exceedances of high thresholds with an application to return level estimation for u.s. wind gusts. *Journal of the American Statistical Association*, 114(528):1865–1879.